

Department of Applied Mathematics
PROBABILITY AND STATISTICS PRELIMINARY EXAMINATION
January 2022

Instructions:

Do two of three problems in each section (Stat and Prob).
Place an **X** on the lines next to the problem numbers
that you are **NOT** submitting for grading.

Prob
1. ____
2. ____
3. ____

Please do not write your name anywhere on this exam.
You will be identified only by your student number.
Write this number **on each page** submitted for grading.
Show all relevant work.

Stat
4. ____
5. ____
6. ____
Total ____

Student Number _____

Probability Section

1. Probability: Problem 1

Consider a random vector (X, Y) taking values in $(0, \infty) \times (0, \infty)$ with the joint probability density function

$$f(x, y) = e^{-(x+y)} [1 + \alpha(2e^{-x} - 1)(2e^{-y} - 1)],$$

where $\alpha \in [-1, 1]$ is a given constant.

- (a) What is the probability density function of Y ?
 - (b) What is the probability density function of X given Y ?
 - (c) Are X and Y independent?
 - (d) Compute the correlation coefficient between X and Y .
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2. Probability: Problem 2

A continuous-time Markov chain X_t is used to model the state of a financial market, which alternates between “*bull*” (the good state) and “*bear*” (the bad state). A statistical analysis shows that “*bull*” turns into “*bear*” with a rate $\lambda > 0$, while “*bear*” turns into “*bull*” with a different rate $\eta > 0$. Suppose that $X_0 = \textit{bull}$.

- (a) Write down the infinitesimal generator (or rate matrix) of the continuous-time Markov chain.
- (b) Let T be the time spent for the market to change to *bear* and go back to *bull*. Find the probability density function of T .

A passively managed mutual fund adjusts its portfolio only when the market state changes. It charges a management fee Z_i in the event of the i^{th} market state change, where $\{Z_i\}_{i \in \mathbb{N}}$ are i.i.d. $\text{Uniform}(0, 100)$ that are independent of the Markov chain X .

- (c) For any $t > 0$, let N_t denote the total number of state changes of the market up to time t and C_t denote the total management fee accumulated up to time t . Show that $\mathbb{E}[C_t] = \mathbb{E}[Z_1]\mathbb{E}[N_t]$.
- (d) For any $t > 0$, compute $\mathbb{E}[C_t \mid X_t = \text{bull}]$.

3. Probability: Problem 3

Let $\{X_i\}_{i \in \mathbb{N}}$ be i.i.d. random variables with $\mathbb{P}(X_i = 1) = p$ and $\mathbb{P}(X_i = -1) = 1 - p$ for some $p \in (0, 1)$. Consider a discrete-time process M defined by

$$M_0 := 0 \quad \text{and} \quad M_t := \sum_{i=1}^t X_i \quad \forall t \in \mathbb{N}.$$

Let τ be the first time M reaches either -1 or 3 .

- (a) If we only focus on the process M up to time τ , we may assume without loss of generality that $M_s := M_\tau$ for $s \geq \tau$. Then, the evolution of M up to time τ can be described using a Markov chain with finite states. Write down the transition matrix P of this Markov chain. Which states are recurrent? Which states are transient?
- (b) Find $\mathbb{E}[\tau]$.
- (c) Your answer in (b) should be a finite number. Hence, we can apply Wald's equation and get $\mathbb{E}[M_\tau] = \mathbb{E}[X_1]\mathbb{E}[\tau]$. From this, find the probability that M reaches 3 before it reaches -1 .
- (d) In the case where $\mathbb{E}[M_\tau] = 0$, we re-scale the process M as follows: for any $n \in \mathbb{N}$, define

$$W_t^{(n)} := \frac{1}{\sqrt{2^n}} M_{2^n t}, \quad \forall t \in \mathcal{D}_n := \left\{ \frac{k}{2^n} : k \in \mathbb{N} \cup \{0\} \right\}.$$

Assume that the limiting process

$$W_t := \lim_{n \rightarrow \infty} W_t^{(n)}, \quad \forall t \in \bigcup_{n \in \mathbb{N}} \mathcal{D}_n = \left\{ \frac{k}{2^m} : k \in \mathbb{N} \cup \{0\}, m \in \mathbb{N} \right\}$$

is well-defined. For any fixed $t \in \bigcup_{n \in \mathbb{N}} \mathcal{D}_n$, find the distribution of W_t .

(Comment: A Brownian motion emerges as the continuous extension of W_t to all $t \geq 0$).

Statistics Section

4. Statistics: Problem 4

Let X_1, X_2, \dots, X_n be a random sample from a $\text{Uniform}(\theta, 2\theta)$ distribution, where $\theta > 0$.

- Find the method of moments (MOM) estimator of θ , $\hat{\theta}_{MOM}$. (Recall that MOM estimators are obtained by equating the sample moments with theoretical moments, and solving for θ).
 - Find the MLE of θ , $\hat{\theta}_{MLE}$, and find a constant k such that $E_\theta(k\hat{\theta}_{MLE}) = \theta$
 - Which of these two estimators can be improved using sufficiency, and how?
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5. Statistics: Problem 5

Let X_1, X_2, \dots, X_n be a random sample from the continuous distribution with probability density function (pdf)

$$f(x; \theta) = \frac{2\theta(1-x)}{(2x-x^2)^{1-\theta}} I_{(0,1)}(x).$$

Here, $\theta > 0$ and $I_{(0,1)}(x)$ is the indicator function that takes on the value 1 when $0 < x < 1$ and is 0 otherwise.

- Find the distribution of $Y_i = -\ln(2X_i - X_i^2)$.
 - Find the maximum likelihood estimator (MLE) for θ . Show that it is an asymptotically unbiased estimator for θ .
 - Find the uniformly minimum variance unbiased estimator (UMVUE) for θ .
 - Is the UMVUE an efficient estimator of θ ? Justify.
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6. Statistics: Problem 6

Let T_1, T_2, \dots, T_n be iid, continuous, non-negative random variables (representing lifetimes for example) from a distribution with pdf $f(t) = f(t; \theta)$ and cdf $F(t) = F(t; \theta)$. Let C_1, C_2, \dots, C_n be iid continuous random variables from a distribution with pdf $g(t)$ and cdf $G(t)$, with fixed, known parameters. Suppose we observe $(X_1, \Delta_1), (X_2, \Delta_2), \dots, (X_n, \Delta_n)$ where

$$X_i = \min(T_i, C_i), \quad \text{for } i = 1, 2, \dots, n,$$

and Δ_i is the indicator random variable taking the value 1 if $T_i \leq C_i$ and is 0 otherwise.

Assume that the X_i and C_i are independent, for each i .

- Write down $h(\vec{X}, \vec{\Delta})$, the joint density of $\{(X_1, \Delta_1), (X_2, \Delta_2), \dots, (X_n, \Delta_n)\}$.

(b) The “hazard function” is defined as:

$$h(t) = \lim_{u \rightarrow 0} \frac{P(t \leq T < t + u | T \geq t)}{u} = \frac{f(t)}{1 - F(t)}.$$

Consider the “joint” hazard function,

$$h(x, \delta) = \lim_{u \rightarrow 0} \frac{P(x \leq X < x + u, \Delta = \delta | X \geq x)}{u}.$$

Give an interpretation of this function specifically when $\delta = 1$.

- (c) Suppose now that the lifetimes T_1, T_2, \dots, T_n are iid exponential random variables with rate λ . Find the MLE (maximum likelihood estimator) of λ based on the observations $(X_1, \Delta_1), (X_2, \Delta_2), \dots, (X_n, \Delta_n)$.
- (d) Estimate the Cramér-Rao lower bound for the variance of all unbiased estimators of λ based on $(X_1, \Delta_1), (X_2, \Delta_2), \dots, (X_n, \Delta_n)$.
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