

Department of Applied Mathematics
PROBABILITY AND STATISTICS PRELIMINARY EXAMINATION
January 2024

Instructions:

Do two of three problems in each section (Prob and Stat).
Place an **X** on the lines next to the problem numbers
that you are **NOT** submitting for grading.

Prob
1. ____
2. ____
3. ____

Do not write your name anywhere on this exam.
You will be identified only by your student number.
Write this number **on each page** submitted for grading.
Show all relevant work!

Stat
4. ____
5. ____
6. ____
Total ____

Student Number _____

Probability Section

Problem 1.

We consider a sequence $\{X_i\}$, $i \geq 1$, of independent random variables such that for each positive integer $i \geq 2$,

$$\text{Prob}(X_i = -i) = \text{Prob}(X_i = i) = \frac{1}{2i \log i}, \quad \text{and} \quad \text{Prob}(X_i = 0) = 1 - \frac{1}{i \log i},$$

with

$$\text{Prob}(X_1 = 0) = 1.$$

Let

$$S_n = \sum_{i=1}^n X_i,$$

(a) (20 points.) Prove that X_i obeys the weak law of large numbers, to wit

$$\forall \varepsilon > 0, \lim_{n \rightarrow \infty} \text{Prob} \left(\left| \frac{S_n}{n} \right| \geq \varepsilon \right) = 0.$$

We will now prove that X_i does not obey the strong law of large numbers. By contradiction, assume that $n^{-1} \sum_{i=1}^n X_i \rightarrow 0$ almost surely.

(b) (30 points.) Let Ω_i be the event $\left| \frac{X_i}{i} \right| \geq 1$. Deduce that

$$\text{Prob} \left(\bigcap_{i_0=1}^{\infty} \bigcup_{i=i_0}^{\infty} \Omega_i \right) = 0.$$

(c) (20 points.) Prove that

$$\lim_{n \rightarrow \infty} \sum_{i=1}^n P(\Omega_i) = \infty.$$

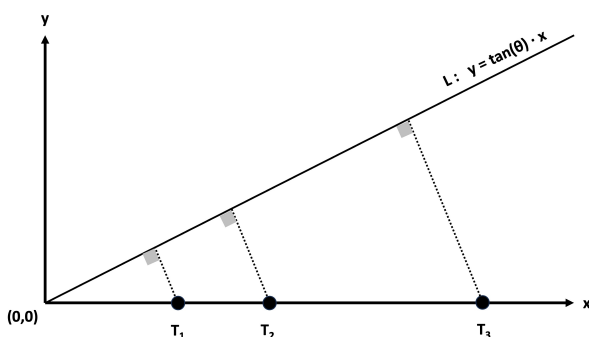
(d) (30 points.) Prove that

$$\text{Prob} \left(\bigcap_{i_0=1}^{\infty} \bigcup_{i=i_0}^{\infty} \Omega_i \right) = 1,$$

and conclude. **Hint.** You may find useful to use that $(1 - x) \leq e^{-x}$, for $x \geq 0$.

Problem 2.

In the xy -plane, consider the arrival times $\{T_n\}_{n \geq 0}$ of a homogeneous Poisson point process (HPPP) with intensity $\lambda > 0$ along the non-negative x -axis. Imagine projecting the process' arrivals orthogonally along the line $L : y = \tan(\theta) \cdot x$, for a given $0 < \theta < \pi/2$ (see the figure below).



Let D_n denote the distance between $(0, 0)$ and the orthogonal projection of T_n along L .

- (30 points.) Determine whether $D := \{D_n\}_{n \geq 1}$ is an HPPP and, if so, its intensity. Justify your answer!
- (10 points.) What's the distribution and expected value of $(D_4 - D_1)$?
- (10 points.) What's the probability of the event $A := [0 < D_1 \leq D_2 \leq D_3 < 1 < D_4]$?
- (10 points.) Given the event A , what's the conditional covariance between D_4 and D_1 ?
- (40 points.) Given the event A , what's the conditional expectation of $(D_4 - D_1)$?

Problem 3.

Consider a time-homogeneous Markov process $X = (X_t)_{t \geq 0}$ with state space

$$S = \{0, 1, 2, \dots\},$$

and rate matrix Q which satisfies $Q(i, j) = 0$ when $|i - j| > 1$.

- (50 points.) Let $\pi : S \rightarrow \mathbb{R}$ be a function with $\pi(i) \geq 0$, for all $i \in S$, and $\sum_{i \in S} \pi(i) = 1$. Show that π is a stationary distribution of X if and only if π satisfies the detailed-balance condition.

- (b) (50 points.) Next, apply the above to address the stationarity of an M/M/3 queue. In particular, X keeps track of the number of customers in line or being served in a queue with an external arrival rate $\lambda > 0$ and three servers, each of which has a service rate $\mu > 0$. Determine a necessary and sufficient condition for this queue to have a stationary distribution, and provide this distribution if the condition is met.
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Statistics Section

Problem 4.

Let X_1, \dots, X_n , with $n \geq 3$, be a random sample from a Poisson distribution with an unknown rate parameter $\lambda > 0$.

- (a) (30 points.) Determine the expected value of $X_1 \cdot I[X_2 = 0]$, where $I[A]$ is the indicator of an event A .
- (b) (70 points.) Determine explicitly the uniformly minimum variance unbiased estimator (UMVUE) of $\lambda e^{-\lambda}$.
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Problem 5.

Let X_1, \dots, X_n be a random sample from the Uniform $[0, \theta]$ distribution, where $\theta > 0$ is unknown.

- (a) (30 points.) Determine a 99% confidence lower-bound for θ based on $X_{(n)} := \max\{X_1, \dots, X_n\}$.
- (b) (30 points.) Show that $X_{(n)}$ is consistent for θ .
- (c) (40 points.) Determine the $\lim_{n \rightarrow \infty} n(\theta - X_{(n)})$ in distribution.
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Problem 6. (100 points.) Let X and Y be two independent random variables uniformly distributed on $[\theta, \theta + 1]$. To decide in favor of the hypothesis $H_0 : \theta = 0$ against the alternative $H_1 : \theta > 0$, we consider two tests defined by

$$\begin{aligned}\Phi_1(X) &: \text{reject } H_0 \text{ if } X > 0.95, \\ \Phi_2(X, Y) &: \text{reject } H_0 \text{ if } X + Y > C.\end{aligned}$$

- (a) (5 points.) Find the value of C so that $\Phi_1(X)$ and $\Phi_2(X, Y)$ have the same size.
- (b) (10 points.) Compute the power function of each test, and sketch both functions on the same figure.
- (c) (5 points.) Prove or disprove that $\Phi_2(X, Y)$ is more powerful than $\Phi_1(X)$.
- (d) (40 points.) Design a test $\Phi_3(X, Y)$ that has the same size as $\Phi_2(X, Y)$, but is more powerful than $\Phi_2(X, Y)$.
- (e) (40 points.) Compute the power of $\Phi_3(X, Y)$.
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